

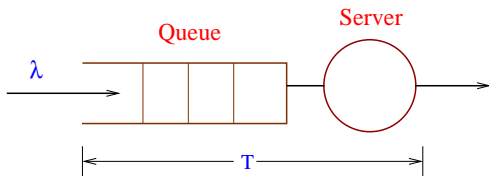
# Poisson processes, Markov chains and M/M/1 queues

Naveen Arulsevan

Advanced Communication Networks

*Lecture 3*

# Little's law



$$N = \lambda T$$

Avg. no. in system

Avg. delay in system

Arrival rate

●  $N$  : Time average / Statistical average

# Little's law applications

## Single server Queue

- Entire system:  $N = \lambda T$
- Just Queue:  $N_Q = \lambda W$
- Just server:  $\rho = \lambda \bar{X}$

## Notation

- $N, N_Q$  = number of customers
- $W, T$  = average delay
- $\lambda$  = arrival rate
- $\bar{X}$  = service time
- $\rho$  = utility

# Single server Queue contd.

## System equations

- Total number in system:  $N = N_Q + \rho$
- Total delay:  $T = W + \bar{X}$
- $N = \lambda T$
- $N_Q = \lambda W$
- $\rho = \lambda \bar{X}$

# Single server Queue contd.

## System equations

- Total number in system:  $N = N_Q + \rho$
  - Total delay:  $T = W + \bar{X}$
  - $N = \lambda T$
  - $N_Q = \lambda W$
  - $\rho = \lambda \bar{X}$
- 
- Fix  $\lambda, \bar{X}$
  - 5 equations, 5 unknowns ( $N, N_Q, W, T, \rho$ )
  - Not independent, Need more info !

# Stochastic Modeling: Arrival times

Easy yet interesting model!

- $\{A(t); t \geq 0\}$  is the arrival process
- $A(t)$  is Number of arrivals in  $[0, t]$
- Basic Model for  $A(t)$ : Poisson Process
- also called *Poisson counting process*  
Non-decreasing, integer valued sample paths

# Poisson process

- $A(t)$  is Poisson process with rate  $\lambda$
- $t > s$  and  $\tau = t - s$
- $A(t) - A(s)$ : Number of arrivals in  $(s, t]$

# Poisson process

- $A(t)$  is Poisson process with rate  $\lambda$
- $t > s$  and  $\tau = t - s$
- $A(t) - A(s)$ : Number of arrivals in  $(s, t]$

## Definition

### Definition 1

- $A(t) - A(s)$  is a Poisson R.V with parameter  $\lambda\tau$  (P1)

$$\Pr(A(t) - A(s) = n) = e^{-\lambda\tau} \frac{(\lambda\tau)^n}{n!} \quad n = 0, 1, 2, \dots$$

- No. of arrivals in any 2 disjoint intervals are independent (P2)

$$\Pr(A(t) - A(s) = n, A(t') - A(s') = n') = \\ \Pr(A(t) - A(s) = n) \cdot \Pr(A(t') - A(s') = n')$$



# Sanity Check

- $A(t)$  is Poisson
- $t_1 < t_2 < t_3$  : 3 time instants
- $B(t_i - t_j)$  denotes  $A(t_i) - A(t_j)$

$$\Pr(B(t_3 - t_1) = 1) = e^{-\lambda(t_3 - t_1)} \lambda(t_3 - t_1) \quad (\text{P1})$$

Alternately,

$$B(t_3 - t_1) = 1 \Rightarrow \begin{cases} B(t_3 - t_2) = 1 \ \& \ B(t_2 - t_1) = 0 \rightarrow \mathbf{E}_1 \\ B(t_3 - t_2) = 0 \ \& \ B(t_2 - t_1) = 1 \rightarrow \mathbf{E}_2 \end{cases}$$

$\mathbf{E}_1$  and  $\mathbf{E}_2$  disjoint

$$\Pr(B(t_3 - t_1) = 1) = \Pr(\mathbf{E}_1) + \Pr(\mathbf{E}_2)$$

## Sanity check Contd..

$$\begin{aligned}\Pr(\mathbf{E}_1) &= \Pr(B(t_3 - t_2) = 1, B(t_2 - t_1) = 0) \\ &= \Pr(B(t_3 - t_2) = 1)\Pr(B(t_2 - t_1) = 0) \quad \text{ind. events, (P2)} \\ &= e^{-\lambda(t_3 - t_2)} \lambda(t_3 - t_2) e^{-\lambda(t_2 - t_1)}\end{aligned}$$

$$\Pr(\mathbf{E}_2) = e^{-\lambda(t_3 - t_2)} e^{-\lambda(t_2 - t_1)} \lambda(t_2 - t_1)$$

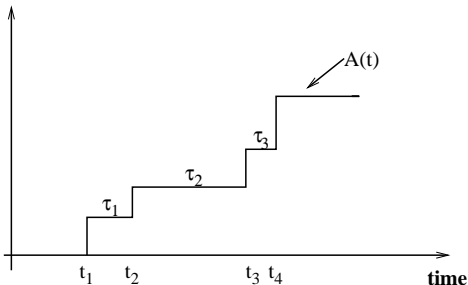
$$\Pr(\mathbf{E}_1) + \Pr(\mathbf{E}_2) = e^{-\lambda(t_3 - t_1)} \lambda(t_3 - t_1)$$

(P1), (P2) consistent for any  $n$

# Poisson: Alternate Viewpoint

Not a counting process?

- $t_n$  = time of  $n$ th arrival  
 $A(t) = n$ ,  $A(t) < n$  for all  $t < t_n$
- $\tau_n = t_{n+1} - t_n$  be the  $n$ th interarrival time
- $A(t)$  determined by sequence of R.Vs  $\tau_1, \tau_2, \dots$



# Poisson Processes

## Definition

Definition 2  $A(t)$  is a Poisson process with rate  $\lambda$  if  $\tau_1, \tau_2, \dots$  are an i.i.d sequence of exponential R.Vs with mean  $\frac{1}{\lambda}$

# Exponential R.Vs

$\tau_n$  is exponential with mean  $\frac{1}{\lambda}$

- CDF:  $\Pr(\tau_n \leq s) = 1 - e^{-\lambda s}, s \geq 0$   
 $0, s < 0$
- $\Pr(\tau_n > s) = e^{-\lambda s}$
- PDF:  $f_{\tau_n}(s) = \lambda e^{-\lambda s}$
- $\text{Var}(\tau_n) = \frac{1}{\lambda^2}$

## Claim

*Defn 1 and Defn 2 are equivalent*

Intuition: Consider zero arrivals in an interval  $[t_n, t_n + s)$

$$\Pr(A(t_n + s) - A(t_n) = 0) = \exp(-\lambda s) \quad (\text{Poisson p.m.f})$$

$$\Pr(\tau_n > s) = \exp(-\lambda s) \quad (\text{Exponential r.v})$$

# Memoryless Property

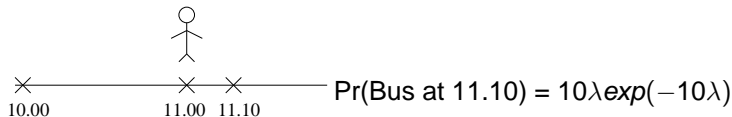
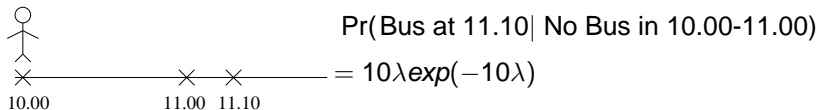
For an exponential R.V  $\tau_n$ ,

$$\Pr(\tau_n > t + s | \tau_n > t) = \Pr(\tau_n > s)$$

$$\begin{aligned}\Pr(\tau_n > t + s | \tau_n > t) &= \frac{\Pr(\tau_n > t + s \text{ and } \tau_n > t)}{\Pr(\tau_n > t)} \\ &= \frac{\Pr(\tau_n > t + s)}{\Pr(\tau_n > t)} = \frac{e^{-\lambda(t+s)}}{e^{-\lambda t}} \\ &= e^{-\lambda s}\end{aligned}$$

Exponentials only continuous R.V with this property

# Example: Bus Arrivals



## Other Useful properties

$A_1(t), A_2(t), \dots, A_k(t)$  are independent Poisson processes with rates  $\lambda_1, \lambda_2, \dots, \lambda_k$ .

$A_{TOT}(t)$  counts total number of arrivals from all processes

$$A_{TOT}(t) = A_1(t) + A_2(t) + \dots + A_k(t)$$

Eg: In network, several input links with Poisson streams combine into one output

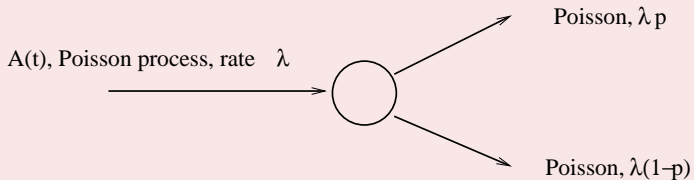
### 1. Adding Independent Poisson processes

$A_{TOT}(t)$  is also Poisson with rate  $\lambda_1 + \lambda_2 + \dots + \lambda_k$



## 2. Splitting

### Randomly split process



Splitting must be independent of arrival times

$A_1(t), A_2(t), \dots, A_N(t)$  are independent *counting* processes  
 $i$ th process has i.i.d interarrival times with mean  $\frac{1}{\mu_i}$  and  
variance  $\sigma_i^2$ .

1. Sum of means:  $\sum_{i=1}^N \frac{1}{\mu_i} = \frac{1}{\lambda}$
2. Sum of variances is finite:  $\sum_{i=1}^N \sigma_i^2 < M$  (constant)

### 3. Limiting Property

As  $N \rightarrow \infty$

$$\sum_{i=1}^N A_i(t) \longrightarrow \text{Poisson}$$

A combination of large number of independent arrivals streams  
can be modeled as Poisson

# Suitability to Network Traffic

- Lots of independent streams in internet
- “Limiting property” - Poisson models reasonable model for network traffic
- Caveat: Sum of variances need to be finite
- Finite variance might not be true

# Service Statistics

- Service time of a packet is Packet size/Link Rate =  $\frac{L}{C}$
- Assume variable length packets  
Service time is exponential with parameter  $\mu$
- $\Pr(X_n \leq x) = 1 - e^{-\mu x}$   
 $\mathbb{E}(X_n) = \frac{1}{\mu} = \frac{\mathbb{E}(L_n)}{C} \triangleq \bar{X}$
- Assume interarrival times and service times are independent

## M/M/1 Queue

FCFS *single server* system, infinite buffer with *Poisson arrivals* and *exponential service times*

Wait till next class to learn more!!